# Estimates of genetic parameters and genotypic values prediction in maizelandrace populations by REML/BLUP procedure 

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#### Abstract

The REML/BLUP statistics are analyses that can be used as selective criteria in the routine of maize breeding programs. The present study aims to determine the genetic potential in crosses of landrace populations applying the REML/BLUP methodology, and to identify populations for the synthesis of new populations and intrapopulation selection for family farming systems, as well as genetic constitutions for use in maize breeding programs. Nine top cross hybrids obtained in the 2012/2013 harvest were evaluated along with their testator, the landraces used as parents, and four commercial hybrids, in a randomized block design, with information taken from the average of each plot. The evaluated traits were: leaf angle, number of ramifications of the tassel, spike insertion height, plant height, spike diameter, number of grains per spike, mass of


grains per spike, spike mass, spike length, prolificity, mass of one hundred grains, and grain yield per plot. The data were analyzed using the SelegenREML/BLUP software. The top cross hybrids Cateto Branco x Planalto, Amarelão x Planalto and the population Cateto Branco are ranked among the ten best crosses, simultaneously, for the traits: leaf angle, number of ramifications of the tassel, spike insertion height, and plant height (Cateto Branco x Planalto), and leaf angle, spike insertion height, and plant height (Amarelão x Planalto and Cateto Branco). The top cross hybrids Criolão x Planalto, Branco 8 Carreiras x Planalto, Caiano Rajado x Planalto, Amarelão x Planalto, Branco Roxo Índio x Planalto stand out for their high genotypic value of the individual BLUP mean components among the ten best genotypes for grain yield, and by combining three or more traits of interest together, being, for effects of selection, the most indicated.

Key words: Zea mays L.; Genetic parameters; Landrace populations; REML/BLUP

## INTRODUCTION

Maize is a crop of international relevance, occupying a prominent position among the agricultural species exploited worldwide, as it is a food of high energetic value and relative low cost, besides being used as raw material in different segments and products, having more than 3500 direct and indirect forms of use. Considering the main cultivated species in the world, maize occupies the third position in terms of cultivated area, but it presents the highest volume of production (Môro and Fritsche-Neto, 2015).

When dealing with the scenario of maize hybrid seeds utilization, conventional or transgenic, along with the use of formulated fertilizers and other agricultural inputs, it must be considered the stratification or establishment of areas of utilization for maize cultivation in the national agriculture. According to available statistical data, the percentage of Brazilian farms that would fit the use of last generation hybrid maize genotypes is very low, among them, the recently released transgenic hybrids. As a result, it is important to implement and constantly search for alternatives within the maize germplasm available in breeding programs and among plant populations traditionally used in agriculture of low employment of modern technological factors (Silveira et al., 2015).

Landrace maize populations, also known as local varieties, are important genetic constituents for genetic breeding programs because of their high adaptive potential for specific environmental conditions (Paterniani et al., 2000). The use of landrace populations present several advantages linked to the sustainability of production, as resistance to diseases, pests and climatic unbalances; also the seeds can be stored and used in the following years, which reduces the cost of production (Carpentieri-Pípolo et al., 2010).

According to Borges et al. (2010) genotype inferences, at any stage of a breeding program, should be based on genotypic and non-phenotypic means, since they reflect future averages when submitted to commercial farming areas.

The consideration of treatment effects as random is significant in genetic plant breeding, because only then the genetic selection can be more efficiently done. Resende (2007b) recommends that, to choose between random and fixed model, the first one should be adopted when the number of treatments in question is more than five, and the fixed model should be used when the number of treatments is smaller.

The mixed models have been frequently used in perennial plants such as eucalyptus (Rosado et al., 2012; Santos et al., 2013), Brazil nut (Camargo et al., 2010), coffee (Carias et al., 2014), sweet potato (Borges et al., 2010), wheat (Pimentel et al., 2014), sugar cane (Bastos et al., 2007), and popcorn (Freitas et al., 2013). However, in the case of maize landrace populations, there is still a small amount of papers using this theory, especially focusing on the evaluation of these populations' potential in crosses.

Therefore, the present study aims to determine the genetic potential in crosses of landrace populations applying the REML/BLUP methodology, and to identify populations for the synthesis of new populations and intrapopulation selection for family farming systems, as well as genetic constitutions for use in maize breeding programs.

## MATERIAL AND METHODS

## Plant material

Seeds of nine maize populations and a variety of open pollination were collected in the municipality of Canguçu, about 50 km from Universidade Federal de Pelotas (UFPel), at the Union of Small Producers of Landraces Corn Seeds, which has been working together with the Center of Genomics and plant breeding (CGF) of UFPel.

The seeds were sown and the crosses were performed in the 2012/2013 growing season between the populations Argentino Branco x BRS Planalto, Dente Ouro x BRS Planalto, Amarelão x BRS Planalto, Criolão x BRS Planalto, Caiano Rajado x BRS Planalto, Branco Oito Carreiras x BRS Planalto, Branco Roxo Índio x BRS Planalto, Cateto Branco x BRS Planalto, and Argentino Amarelo x BRS Planalto, originating nine top crosses. In the agricultural crop 2013/2014 the populations were sown along with the top crosses, other four commercial hybrids were sow: Coodetec 308, Pioneer 30F53, Agroeste 1551 VTPRO, and Agroeste 1590, and also all populations involved in the crosses, resulting in 22 treatments.

## Experimental design

The experiment was conducted at the experimental area of the Center of Genomics and plant breeding of UFPel, city of Capão do Leão, with geographic coordinates of $31^{\circ} 52^{\prime} 00^{\prime \prime} \mathrm{S}$ and $52^{\circ} 21^{\prime} 24^{\prime \prime} \mathrm{O}$, and 13 m in altitude. The soil was classified as dystrophic yellow red Argisol (Santos et al., 2006).

It was used a complete randomized block design with three repetitions. The experimental plots consisted of two rows of 5 m in length, with 42 plants per plot, spaced 0.7 m between rows, corresponding to a sowing density of 60,000 plants $/ \mathrm{ha}$.

## Assessed traits

The following agronomic traits were evaluated in each experimental plot: 1) leaf angle (LA), it refers to the measurement of the first leaf just below the first ear in three plants, given in degrees; 2) number of branches of the tassel (NBT): it refers to the total number of branches of the tassel, given in units; 3) ear insertion height (EIH): it refers to the height from the soil to the insertion of the first ear, given in $\mathrm{cm} ; 4)$ plant height $(\mathrm{PH})$ : it refers to the height from the soil to the insertion of the node of the last leaf, given in $\mathrm{cm} ; 5$ ) ear diameter (ED):
it refers to the average diameter of the central part of three ears, given in mm; 6) number of kernels per ear (NKE): it refers to the average of the total number of kernels in three ears, given in units; 7) kernel mass per ear (KME): it refers to the average of total kernel weight of three ears, obtained by precision scale, given in g; 8) ear mass (EM): it refers to the average weight of three ears before processing, obtained by precision scale, given in $g$; 9) ear length (EL): it refers to the average of total length of three ears, given in cm; 10) prolificity (PRO): it refers to the ratio between the total number of ears by the total number plants in the plot, given in units; 11) one hundred-kernel mass (HKM): it refers to the average of eight assessments of hundred-kernel mass each plot, given in g ; 12) grain yield per plot (GY): it refers to the total grain yield, adjusted to 42 plants per plot and $13 \%$ moisture, given in kg per plot.

## Statistical analysis

The variance components and genetic parameters were obtained using the SelegenREML/BLUP software (Resende, 2007a), model 21 (randomized blocks, autogamous or hybrid test, plot mean) denoted in matrix form by:

$$
\begin{equation*}
\mathrm{y}=\mathrm{Xr}+\mathrm{Zg}+\mathrm{e} \tag{Equation1}
\end{equation*}
$$

where $y$ is the data vector, $r$ is the vector of replications effects (assumed to be fixed) plus the general mean, g is the vector of genotypic effects (assumed to be random), and $e$ is the vector of errors (random), and X and Z, matrices of incidence for these effects (Resende, 2002).

The deviances were obtained by analyzing the model with and without the $\mathrm{h}^{2}$ values for each trait, and then subtracting and confronting them with the chi-square value with 1 and $5 \%$ degrees of freedom of probability, respectively (Resende, 2007a).

## RESULTS AND DISCUSSION

Table 1 shows the deviance values for the studied traits. These results demonstrate the existence of a significant difference by the chi-square test at $5 \%$ of probability, for the genotype effect, as well as its components of variance $\left(\hat{\sigma}^{2}\right)$ and determination coefficients $\left(\hat{h}_{g}^{2}\right)$ of the respective effect. Therefore, the deviance anălysis evidenced the presence of genetic difference among the genotypes analyzed for all evaluated traits.

The results concerning genetic parameters estimates for the traits on study are presented in Table 2. The coefficient of genetic variation $\left(C V_{g}\right)$ is a parameter commonly used to compare the genetic variability evidenced for each trait, as quoted by Resende (2002). The coefficients of genetic variation $\left(C V_{g i}\right)$ were moderate to high for all analyzed traits, ranging between $7.276 \%$ for ear diameter and $37.739 \%$ for grain yield per plot, allowing to infer that the evaluated genotypes present high genetic variability. The values for the coefficient of residual variation $\left(C V_{e}\right)$ were considered mean to low, with values ranging from $12.336 \%$ (prolificity) to $3.349 \%$ (ear diameter).

The magnitude of $C V_{e}$ is associated with the experimental accuracy and way of measurement of the set of analyzed traits. The prolificity, in general, presented higher values of this coefficient, because it is a trait obtained by two other traits, accumulating the measurement errors of these two. However, the coefficients of variation of all traits are within acceptable limits for agricultural experimentation, according to Scapim et al. (1995) and Pimentel-Gomes
(2009). According to Vencovsky (1987), when a relation between $C V_{g i}$ and coefficient of residual variation $C V$ of 1 or more is evidenced, a favorable situation to achieve gains in the selection occurs. Therefore, considering the relation between $C V_{g i}$ and $C V_{e}$ evidenced by $C V_{r}$ for all traits, $C V_{g i}$ was superior to $C V_{e}$, which indicates favorable gains with the selection, highlighting the trait grain yield, with the highest magnitude for this coefficient.

Table 1. Deviance values for nine landrace maize top cross hybrids with the tester BRS Planalto, eight genitors and four commercial witnesses for the traits leaf angle (LA), number of branches of the tassel (NBT), ear insertion height (EIH), plant height (PH), ear diameter (ED), number of kernels per ear (NKE), kernel mass per ear (KME), ear mass (EM), ear length (EL), prolificity (PRO), one hundred-kernel mass (HKM), and grain yield per plot (GY), in the 2013/2014 agricultural year.

| LA | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Genotypes | 301.10 | 40.88** | $\hat{\sigma}_{g}^{2}=28.80^{* *}$ | $\hat{h}_{g}^{2}=0.7621^{* *}$ |
|  | Complete Model | 260.22 |  |  |  |
| NBT | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 255.83 | 77.17** | $\hat{\sigma}_{g}^{2}=16.67^{* *}$ | $\hat{h}_{g}^{2}=0.905^{* *}$ |
|  | Complete Model | 178.66 |  |  |  |
| EIH | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 439.86 | 53.69** | $\hat{\sigma}_{g}^{2}=283.58^{* *}$ | $\hat{h}_{g}^{2}=0.829^{* *}$ |
|  | Complete Model | 386.17 |  |  |  |
| PH | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 485.56 | 48.35** | $\hat{\sigma}_{g}^{2}=568.02^{* *}$ | $\hat{h}_{g}^{2}=0.804^{* *}$ |
|  | Complete Model | 437.21 |  |  |  |
| ED | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 232.98 | 52.77** | $\hat{\sigma}_{g}^{2}=10.58^{* *}$ | $\hat{h}_{g}^{2}=0.825^{* *}$ |
|  | Complete Model | 180.21 |  |  |  |
| NKE | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 661.00 | 46.11** | $\hat{\sigma}_{g}^{2}=9,066.98^{* *}$ | $\hat{h}_{g}^{2}=0.792^{* *}$ |
|  | Complete Model | 614.89 |  |  |  |
| KME | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 654.63 | 82.69** | $\hat{\sigma}_{g=9,482.54^{* *}}^{2}$ | $\hat{h}_{g}^{2}=0.917^{* *}$ |
|  | Complete Model | 571.94 |  |  |  |
| EM | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 662.31 | 66.92** | $\hat{\sigma}_{g}^{2}=10,250.90^{* *}$ | $\hat{h}_{g}^{2}=0.877^{* *}$ |
|  | Complete Model | 595.39 |  |  |  |
| EL | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 155.72 | 27.22** | $\hat{\sigma}_{g=2.47^{* *}}^{2}$ | $\hat{h}_{g}^{2}=0.656^{* *}$ |
|  | Complete Model | 128.5 |  |  |  |
| PRO | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | -137.02 | 17.54** | $\hat{\sigma}_{g=0.019^{* *}}^{2}$ | $\hat{h}_{g}^{2}=0.544^{* *}$ |
|  | Complete Model | -154.56 |  |  |  |
| HKM | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 275.49 | 46.63** | $\hat{\sigma}_{g=20.02^{* *}}^{2}$ | $\hat{h}_{g}^{2}=0.795^{* *}$ |
|  | Complete Model | 228.86 |  |  |  |
| GY | Effect | Deviance | LRT | Var. Comp. | Determ. Coef. |
|  | Genotypes | 158.46 | 86.48** | $\hat{\sigma}_{g=3.63^{* *}}^{2}$ | $\hat{h}_{g}^{2}=0.924^{* *}$ |
|  | Complete Model | 71.98 |  |  |  |

*Significant at $5 \%$ of probability of error by the chi-square test. **Significant at $1 \%$ of probability of error by the chi-square test for the variables, and ${ }^{\text {ns }}$ non-significant.

Table 2. Genetic parameters estimates of nine landrace maize top cross hybrids, tester BRS Planalto, eight genitors and four commercial witnesses for the traits leaf angle (LA), number of branches of the tassel (NBT), ear insertion height (EIH), plant height (PH), ear diameter (ED), number of kernels per ear (NKE), kernel mass per ear (KME), ear mass (EM), ear length (EL), prolificity (PRO), hundred-kernel mass (HKM), and grain yield per plot (GY), in the 2013/2014 agricultural year.

| Estimates | LA | MHG | NRT | SIH | PH | SD | NGS | SM | MGS | PRO | SL | GY |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\hat{\sigma}_{g}^{2}$ | 28.805 | 20.019 | 16.674 | 283.582 | 568.022 | 10.577 | $9,066.981$ | $10,250.901$ | $9,482.539$ | 0.020 | 2.467 | 3.631 |
| $\hat{\sigma}_{e}^{2}$ | 8.989 | 5.153 | 1.750 | 58.406 | 138.322 | 2.241 | $2,373.475$ | $1,430.093$ | 857.621 | 0.016 | 1.294 | 0.297 |
| $\hat{\sigma}_{p}^{2}$ | 37.794 | 25.172 | 18.424 | 341.987 | 706.344 | 12.819 | $11,440.45$ | $11,680.994$ | $10,340.160$ | 0.036 | 3.760 | 3.928 |
| $\hat{h}_{g}^{2}$ | 0.762 | 0.795 |  |  |  |  |  |  |  |  |  |  |
| $\pm 0.304$ | 0.310 | 0.905 |  |  |  |  |  |  |  |  |  |  |
| $\pm 0.331$ | 0.829 |  |  |  |  |  |  |  |  |  |  |  |
| $\pm 0.317$ | 0.804 <br> $\pm 0.312$ | 0.825 <br> $\pm 0.316$ | 0.792 <br> $\pm 0.310$ | 0.877571 <br> $\pm 0.3261$ | 0.917 <br> $\pm 0.333$ | 0.544 <br> $\pm 0.257$ | 0.656 <br> $\pm 0.282$ | 0.924 <br> $\pm 0.334$ <br> $\hat{h}_{m c}^{2}$ | 0.906 | 0.921 | 0.966 | 0.936 |
| 0.925 | 0.934 | 0.920 | 0.955 | 0.971 | 0.782 | 0.851 | 0.973 |  |  |  |  |  |
| Aclinh | 0.952 | 0.960 | 0.983 | 0.967 | 0.962 | 0.966 | 0.959 | 0.977 | 0.985 | 0.884 | 0.923 | 0.987 |
| $\mathrm{CV}_{\mathrm{gi}}(\%)$ | 19.206 | 11.153 | 26.504 | 14.039 | 10.611 | 7.276 | 22.005 | 17.272 | 19.706 | 13.482 | 8.621 | 37.739 |
| $\mathrm{CV}_{\mathrm{c}}(\%)$ | 10.729 | 5.658 | 8.587 | 6.371 | 5.236 | 3.349 | 11.259 | 6.451 | 5.926 | 12.336 | 6.243 | 10.792 |
| CV F | 1.790 | 1.971 | 3.087 | 2.203 | 2.026 | 2.172 | 1.952 | 2.677 | 3.325 | 1.093 | 1.381 | 3.497 |
| GA | 27.945 | 40.118 | 15.407 | 119.947 | 224.609 | 44.699 | 432.720 | 586.175 | 494.151 | 1.039 | 18.219 | 5.049 |

$\hat{\sigma}^{2}=$ genotypic variance; $\hat{\sigma}^{2}=$ residual variance; $\hat{\sigma}_{p}^{2}=$ individual phenotypic variance; $\hat{h}_{g}^{2}=$ heritability in the broad sense of individual plots; $h_{m c}^{2}=$ heritability of the genotype mean; Aclinh $=$ genotypes selective accuracy; $\mathrm{CV}_{\mathrm{gi}}(\%)=$ coefficient of genotypic variance; $\mathrm{CV}_{\mathrm{e}}(\%)=$ coefficient of residual variation; $\mathrm{CV}_{\mathrm{r}}=$ coefficient of relative variation; $G A=$ general average.

The heritability estimates revealed the reliability of the phenotypic value as indicator of reliability. For this reason, heritability almost always participates in expressions related to the prediction of gains of different breeding methods and, as a consequence, of the decisions that breeders make conducting their selection programs (Ramalho et al., 2012). The coefficient of individual heritability in the broad sense $\left(h_{g}^{2}\right)$ extracts the additive and dominance variance between the selection units of the phenotypic variance, which serves the purpose of hybrid selection, since the entire genetic variance is exploited in the F1 generation. High heritability values were found for all traits, with grain yield (0.924), mass of grains per spike (0.917) and number of ramifications of the tassel ( 0.905 ) with the highest magnitudes. These results reveal a favorable condition for genetic gains with selection using the considered traits. According to the classification described by Resende (2002), heritability can be considered of low magnitude when $\hat{h}_{a}^{2}<0.15$, mean magnitude between $0.15<\hat{h}_{a}^{2}<0.50$, and high magnitude with $\hat{h}_{a}^{2}<$ 0.50. Therefore, it can be affirmed that all analyzed traits showed high heritability. According to Batista et al. (2012) the value of heritability may be affected if there is a change in any of the variance components. Thus, these high heritability coefficients indicated significant genetic control on the evaluated traits. Therefore, it is possible to infer that the traits studied in this set of maize genotypes may be used to evaluate genetic variability and are efficient for selection prediction. Estimates based on the mean of genotypes $\left(h_{m c}^{2}\right)$ were high for all traits, with amplitudes of 0.772 and 0.973 for prolificity and grain yield, respectively. These results presented higher magnitudes than heritability at individual level $\left(h_{g}^{2}\right)$ for all traits. The heritability of the genotypes average is estimated when using means as an evaluation/selection parameter. According to Maia et al. (2009), this is the quotient of interest to aim for breeding success, once the genotypes selection considers their predicted genotypic values, based on means of several repetitions.

The selective accuracy (Aclinh) reflects the quality of results and procedures used to predict genetic values. This parameter is associated to the accuracy of selection and refers to the correlation between predicted genetic values and true genetic values of individuals (Pimentel et al., 2014). The results of selective accuracy presented an amplitude of 0.884 for prolificity and 0.987 for grain yield, being considered of high magnitude. This statistic ranges from 0 to 1 and, according to Resende and Duarte (2007), it may be classified as very high (Aclinh $\geq 0.90$ ), high $(0.70 \leq$ Aclinh $<0.90)$, moderate $(0.50 \leq$ Aclinh $<0.70)$ and low (Aclinh $<0.50$ ). Values of accuracy higher than 0.70 are sufficient to provide a precise and efficient inference about the genetic value of genotypes, and as it is a measure associated with precision in selection, accuracy is the main element of genetic progress that may be altered by the breeder, allowing to maximize genetic gains (Resende, 2002).

Tables 3, 4 and 5 show the individual BLUP mean components in nine maize top cross hybrids of landraces, BRS Planalto tester, eight genitors, and four commercial hybrids used as control treatment. The genotype values of each genotype were obtained by adding each genotypic effect to the overall mean of the experiment. The genetic gain was equivalent to the mean of predicted genetic effects vectors for the selected genotypes. The overall mean plus the genetic gain results in the improved population mean.

Analyzing the ten best genotypes by order, five (Cateto Branco x Planalto, Agroceres 1590, Coodetec 308, Agroeste 1551 VT PRO 2, and Pioneer 30F53H) of them concomitantly presented the lowest genotypic values for the traits: leaf angle, number of ramifications of the tassel, spike insertion height and plant height; two (Amarelão x Planalto and Cateto Branco) presented the lowest values for leaf angle, spike insertion height and plant height; one (Argentino Branco) showed the smallest magnitudes for leaf angle, number of ramifications of the tassel and spike insertion height, and one (Branco Oito Carreiras) stood out for number of branches of the tassel, spike insertion height and plant height. It was also verified that genotype Argentino Amarelo presented the best performances for number of ramifications of the tassel and spike insertion height. The top cross hybrids Cateto Branco x Planalto, Amarelão x Planalto and the population Cateto Branco evidenced, at the same time, among the ten superior genotypes in four, three and three traits, respectively. These results show the possibility of finding, within these genetic constitutions, individuals that concentrate favorable alleles for this set of traits simultaneously (Tables 3 and 4).

The plant size reduction is one goal of selection, according to Jesus Freitas et al. (2013), plants with a higher size and higher spike insertion difficult mechanical harvesting, besides being more susceptible to breaking and lodging, especially when it comes to regions with a high incidence of winds.

Regarding some tassel traits, Duvick and Cassman (1999) relate an important modification made by breeding programs in their group of hybrids grown in the North American corn belt, selecting for tassel size reduction. According to the authors, smaller tassels cause less apical dominance on the spikes, which is very important under stress conditions. In addition, it is possible that the lower expenses by the plant for developing the male inflorescence contributes to a better adaptation on higher plant densities per unit of area. Sangoi et al. (2006), based on the results of an experiment combining three factors: plant population, hybrids and detasseling levels, concluded that tassel removal significantly increased grain yield. The authors found an increase of $7 \%$ in average yield of the tested genotypes submitted to detasseling.
Table 3. Ordering (Ord), genotypes (Gen), predicted genotypic effect (g), genotypic values ( $\mathrm{g}+\mu$ ), predicted genetic gains (Gain), and new population mean (New
mean) in nine landrace maize top cross hybrids, tester BRS Planalto, eight genitors and four commercial witnesses for the traits leaf angle (LA), number of branches
of the tassel (NBT), ear insertion height (EIH), hundred-kernel mass (HKM), in the $2013 / 2014$ agricultural year. CGF/FAEM-UFPel, Pelotas, RS, Brazil.

${ }^{1} 1$ = Argentino Branco x Planalto, $2=$ Branco Roxo Índio x Planalto, $3=$ Amarelão x Planalto, $4=$ Caiano Rajado x Planalto, 5 =Argentino Amarelo x Planalto, 6 $=$ Dente de Ouro x Planalto, $7=$ Cateto Branco $\times$ Planalto, $8=$ Branco 8 Carreiras $\times$ Planalto, $9=$ Criolão $\times$ Planalto, $10=$ Agroceres $1590,11=$ Coodetec 308, 12 $=$ Agroeste 1551 VT PRO 2, $13=$ Pioneer 30F53H, $14=$ Planalto, $15=$ Argentino Branco, $16=$ Dente Ouro, $17=$ Amarelão, $18=$ Criolão, $19=$ Caiano Rajado,
$20=$ Branco Oito Carreiras, $21=$ Cateto Branco, and $22=$ Argentino Amarelo.

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Table 4. Ordering (Ord), genotypes (gen), predicted genotypic effect ( g ), genotypic values ( $\mathrm{g}+\mathrm{u}$ ), predicted genetic gains and new population mean in
nine landrace maize top cross hybrids, tester BRS Planalto, eight genitors and four commercial witnesses for the traits plant height (PH), ear diameter (ED),
number of kernels per ear (NKE) and kernel mass per ear (KME) in the 2013/2014 agricultural year. CGF/FAEM-UFPel, Pelotas, RS, Brazil.

${ }^{1} 1=$ Argentino Branco x Planalto, $2=$ Branco Roxo Índio x Planalto, $3=$ Amarelão x Planalto, $4=$ Caiano Rajado x Planalto, $5=$ Argentino Amarelo x Planalto, 6 = $20=$ Branco Oito Carreiras, $21=$ Cateto Branco, and $22=$ Argentino Amarelo.

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$1=$ Argentino Branco x Planalto, $2=$ Branco Roxo Índio x Planalto, $3=$ Amarelão x Planalto, $4=$ Caiano Rajado x Planalto, $5=$ Argentino Amarelo x Planalto, 6 $=$ Dente de Ouro x Planalto, $7=$ Cateto Branco x Planalto, $8=$ Branco 8 Carreiras x Planalto, $9=$ Criolão x Planalto, $10=$ Agroceres $1590,11=$ Coodetec 308, 12 $=$ Agroeste 1551 VT PRO 2, $13=$ Pioneer 30 F53H, $14=$ Planalto, $15=$ Argentino Branco, $16=$ Dente Ouro, $17=$ Amarelão, $18=$ Criolão, $19=$ Caiano Rajado,
$20=$ Branco Oito Carreiras, $21=$ Cateto Branco, and $22=$ Argentino Amarelo.

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For grain yield, the genotypes Agroeste 1551 VT PRO 2, Pioneer 30F53H, Agroceres 1590, Coodetec 308, Planalto, and the crosses Criolão x Planalto, Branco 8 Carreiras x Planalto, Caiano Rajado x Planalto, Branco Roxo Índio x Planalto, and Amarelão x Planalto presented the ten best performances by ordering the genotypic values (Table 5). The top cross hybrid Criolão x Planalto presented, concomitantly, among the top ten for number of grains, mass of grainsper spike, spike mass, spike length; the top cross Branco 8 Carreiras x Planalto, for spike diameter, number of grains, mass of grains per spike, spike mass, spike length, and mass of one hundred grains; the top cross Caiano Rajado x Planalto, for spike diameter, spike mass, prolificity, spike length, and mass of one hundred grains; the top cross Branco Roxo Índio x Planalto, for spike diameter, mass of grains per spike, spike mass, and mass of one hundred grains; and the top cross 3 for spike diameter, mass of grains per spike, spike length, and mass of one hundred grains (Tables 3, 4 and 5).

The results show that these top cross hybrids may subsidize the establishment of a genetic breeding program for the species, even presenting inferior performance than the commercial materials regarding grain yield. These genetic constitutions are important because they are a source of genetic variability that can be exploited in the search for genes tolerant and/or resistant to biotic and abiotic factors as reported by Araújo and Nass (2002).

In general, the grain yield of the top cross hybrids and landraces were smaller than the commercial hybrids. However, according to Carpentiere-Pípolo et al. (2010), although they are considered to be less productive than commercial genotypes, they present high productive potential under low technology cultivation conditions. These authors also add that those populations, which have their productivities between 50 and $70 \%$ in relation to commercial witnesses, may be useful for breeding if they present other traits of interest.

According to Arnhold et al. (2012), analyzing 128 different maize genotypes in two years, reported a correlation of 0.75 between phenotypic and genotypic means using mixed models, showing that the selection based on phenotypic means may not reflect the selection of the same individuals by genotypic means. The same authors also reported that, in occasion of data loss, the mixed models are preferable. Freitas et al. (2013) evaluated progenies of popcorn half-siblings using the REML/BLUP method, which were highly efficient when confronted with several selection indexes, making it possible to select progenies with high relative performances and promising predicted genetic gains.

Due to the greater variability $\left(\mathrm{CV}_{\mathrm{gi}}\right)$ and heritability coefficient of the trait grain yield in the present study, the genotypic values, as well as their inferior (ILCI) and superior (SLCI) limits of the $95 \%$ confidence interval were obtained, presented in Figure 1. It was verified that the commercial hybrid Agroeste 1551 VT PRO 2 differed from the other genotypes, based on the confidence intervals overlap, in relation to their predicted genotypic values. It was also verified that the top cross Criolão x Planalto, which presented the best ranking among the other top cross hybrids by their genotype value (Table 5), had inferior performance only than the commercial hybrid Agroeste 1551 VT PRO 2, being a promising genotype, although revealing no differences of genotypic effects from the nine best genotypes by the ordering. Similarly, when considering the confidence intervals overlapping of genotypic values, we observed that several genotypes are statistically equal at $95 \%$ of confidence level, being unable to prove differences between them, as reported by Resende (2006).

The coefficient of genetic variation presents a major contribution to the total phenotypic variance for all evaluated traits. The individual heritabilities in the broad sense, and based on genotypes means, contribute to greater genetic gains regarding the grain yield trait.


Figure 1. Predicted genotypic values $(\hat{\mu}+\hat{g})$ and their confidence intervals $(95 \%)$ for the grain yield (GY) trait in nine landrace maize top cross hybrids, tester BRS Planalto, eight genitors and four commercial witnesses.

The top cross hybrids Cateto Branco x Planalto, Amarelão x Planalto and the population Cateto Branco are the best classified among the ten best genotypes, simultaneously, for the traits: leaf angle, number of ramifications of the tassel, spike insertion height and plant height, and leaf angle, spike insertion height and plant height (Amarelão x Planalto and Cateto Branco).

The top cross hybrids Criolão x Planalto, Branco 8 Carreiras x Planalto, Caiano Rajado x Planalto, Amarelão x Planalto, and Branco Roxo Índio x Planalto stand out for presenting high genotypic value of individual BLUP mean components among the ten best genotypes for grain yield, and combining together three or more traits of interest, being, for purpose of selection, the most indicated.

## Conflicts of interest

The authors declare no conflict of interest.

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